

Some Consequences of a Symmetry in Strong Distributions*

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Some polynomials and interpolatory quadrature rules associated with strong Stieltjes distributions are considered, especially when the distributions satisfy a certain symmetric property. © 1995 Academic Press, Inc.

1. INTRODUCTION

Let the function $\psi(t)$, defined on (a, b) , be real, bounded, and nondecreasing with infinitely many points of increase in (a, b) , and such that the moments

$$\mu_m = \int_a^b t^m d\psi(t), \quad m = 0, \pm 1, \pm 2, \dots,$$

all exist. Then $d\psi(t)$ is called a strong distribution on (a, b) .

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We refer to a strong distribution on (a, b) as a strong Stieltjes distribution on (a, b) , or simply as a $SS(a, b)$ distribution, when $(a, b) \subseteq (0, \infty)$. Given a $SS(a, b)$ distribution $d\psi(t)$ it is known, see for example [6], that

$$\int_a^b t^{-n+s} B_n(t) d\psi(t) = 0, \quad 0 \leq s \leq n-1, n \geq 1, \quad (1.1)$$

defines a unique set of polynomials $B_n(z)$, taken to be monic, satisfying the three-term recurrence relation

$$B_n(z) = (z - \beta_n)B_{n-1}(z) - \alpha_n z B_{n-2}(z), \quad n \geq 2, \quad (1.2)$$

with $B_0(z) = 1$, $B_1(z) = z - \beta_1$, and $\beta_n > 0$, $\alpha_{n+1} > 0$ for $n \geq 1$. For these results and others, but in terms of the equivalent orthogonal Laurent polynomials, see [4].

In [6], the polynomials $B_n(z)$ and the associated interpolatory quadrature rules have been investigated for a class of $SS(a, b)$ distributions, denoted by $ScS(a, b)$ distributions, which satisfy the symmetric (inverse) property

$$\frac{d\psi(t)}{\sqrt{t}} = -\frac{d\psi(c/t)}{\sqrt{c/t}}, \quad t \in (a, b).$$

In this article we look at the real polynomials $B_n(\lambda, z)$ given by

$$B_n(\lambda, z) = B_n(z) - \lambda B_{n-1}(z) \quad n \geq 1, \quad (1.3)$$

where $\lambda \in \mathbb{R}$, and the associated interpolatory quadrature rules for a class of $SS(a, b)$ distributions which possess the symmetric (inverse) property

$$d\psi(t) = -d\psi(c/t), \quad t \in (a, b), \quad (1.4)$$

for $c > 0$. Just as in the $ScS(a, b)$ distributions we must have $a = 0$ iff $b = \infty$ and, if $0 < a < b < \infty$ then $c = ab$. When $d\psi(t)$ can be given in the form $w(t) dt$ then $w(t)$ satisfies

$$t w(t) = (c/t) w(c/t), \quad t \in (a, b).$$

For convenience we denote the class of distributions which satisfy (1.4) as $S\bar{C}S(a, b)$ distributions.

LEMMA 1.1. *Let $d\psi(t)$ be a $S\bar{C}S(a, b)$ distribution and let $f(t)$ be integrable with respect to $d\psi(t)$ on (a, b) . Then*

$$\int_a^b f(t) d\psi(t) = \int_a^b f(c/t) d\psi(t).$$

The proof of this lemma follows from change of variable properties of Riemann–Stieltjes integrals and from the choice of c .

Some properties of the special $\text{S}\bar{\text{C}}\text{S}(a, b)$ distributions, where $c = 1$, have been considered in [3]. The $\text{S}\bar{\text{C}}\text{S}(a, b)$ distributions taking the form $w(t) dt$ have been studied in [8]. It was shown there that for any such distributions

$$\begin{aligned} \frac{z^n B_n(c/z)}{B_n(0)} &= B_n(\alpha_{n+1}, z) \\ &= B_n(z) - \alpha_{n+1} B_{n-1}(z) \\ &= \{B_{n+1}(z) + \beta_{n+1} B_n(z)\}/z, \quad n \geq 1, \end{aligned} \quad (1.5)$$

and, further, if $\gamma_n = \beta_n + \alpha_{n+1}$, then

$$\frac{\gamma_{n+1}}{\gamma_n} = \frac{c}{\beta_n \beta_{n+1}}, \quad n \geq 1, \quad (1.6)$$

with $\gamma_1 = c/\beta_1$. These results can easily be proved for all $\text{S}\bar{\text{C}}\text{S}(a, b)$ distributions.

For a given $\text{S}\bar{\text{C}}\text{S}(a, b)$ distribution $d\psi(t)$ it follows from Lemma 1.1 that $\mu_m = c^m \mu_{-m}$, $m \geq 0$. Consequently for the associated Hankel determinants

$$H_0^{(m)} = 1, \quad H_n^{(m)} = \begin{bmatrix} \mu_m & \mu_{m+1} & \cdots & \mu_{m+n-1} \\ \mu_{m+1} & \mu_{m+2} & \cdots & \mu_{m+n} \\ \vdots & \vdots & & \vdots \\ \mu_{m+n-1} & \mu_{m+n} & \cdots & \mu_{m+2n-2} \end{bmatrix}, \quad n \geq 1,$$

$m = 0, \pm 1, 2, \dots$, the following holds:

$$H_n^{(m)} = c^{n(m+n-1)} H_n^{(-m-2n+2)}, \quad m \geq 0, n \geq 0.$$

2. THE POLYNOMIALS $B_n(\lambda, z)$

In this section $d\psi(t)$ is assumed to be any $\text{SS}(a, b)$ distribution. In (1.3) the polynomials $B_n(\lambda, z)$ are defined by $B_n(z)$ in just the same way as the quasi-orthogonal polynomials (see, for example, [1]) are defined from orthogonal polynomials.

It follows from (1.1) and (1.3) that

$$\int_a^b t^{-n+s} B_n(\lambda, t) d\psi(t) = 0, \quad 1 \leq s \leq n-1, n \geq 2. \quad (2.1)$$

A more general result is the following.

THEOREM 2.1. *Let $Q_n(z)$ be a real monic polynomial of degree $n \geq 2$ which satisfies the condition*

$$\int_a^b t^{-n+s} Q_n(t) d\psi(t) = 0, \quad 1 \leq s \leq n-1.$$

Then there exists a $\lambda \in \Re$ such that $Q_n(z) = B_n(\lambda, z)$.

This result is easily proved by considering the linear combination of $Q_n(t)$ in terms of the polynomials $B_r(t)$, $r = 1, 2, \dots, n$.

It can be established from (2.1) that the zeros of $B_n(\lambda, z)$ are all real and distinct and that at least $n-1$ of them lie inside (a, b) . Let these zeros, in increasing order, be denoted by $z_{n,1}^{(\lambda)}, z_{n,2}^{(\lambda)}, \dots, z_{n,n}^{(\lambda)}$. From (1.2) and (1.3), $B_n(\lambda, 0) = (-1)^n \beta_1 \dots \beta_{n-1} (\beta_n + \lambda)$ and, hence, for example, we can say that if $\lambda > -\beta_n$ then $z_{n,1}^{(\lambda)} > 0$ while if $\lambda < -\beta_n$ then $z_{n,1}^{(\lambda)} < 0$.

THEOREM 2.2. *For $n \geq 1$, if $0 \leq \lambda \leq \alpha_{n+1}$ then all the zeros of $B_n(\lambda, z)$ lie inside the interval (a, b) .*

Proof. It is known that the zeros of $B_n(0, z) = B_n(z)$ and $B_n(\alpha_{n+1}, z)$ lie inside (a, b) . Thus, we obtain

$$\begin{aligned} (-1)^n B_n(0, a) &> 0, & (-1)^n B_n(\alpha_{n+1}, a) &> 0, & B_n(0, b) &> 0, \\ & & & & B_n(\alpha_{n+1}, b) &> 0 \end{aligned}$$

for all $n \geq 1$. From this and from

$$B_n(\lambda, z) = \left(1 - \frac{\lambda}{\alpha_{n+1}}\right) B_n(0, z) + \frac{\lambda}{\alpha_{n+1}} B_n(\alpha_{n+1}, z),$$

we then have for $0 \leq \lambda \leq \alpha_{n+1}$

$$(-1)^n B_n(\lambda, a) > 0, \quad B_n(\lambda, b) > 0, \quad n \geq 1. \quad (2.2)$$

It is true that $z_{n,n-1}^{(\lambda)} < b$. Now, if $z_{n,n}^{(\lambda)} > b$ then we must have $B_n(\lambda, b) < 0$, which is a contradiction to (2.2). Hence, $z_{n,n}^{(\lambda)} < b$.

Similarly, we can see that $z_{n,1}^{(\lambda)} < a$ also leads to a contradiction of (2.2). Hence the theorem is proved.

From (1.2) and (1.3) it can be seen that when $n \geq 2$, the zeros of $B_n(\lambda, z)$ for any λ , are also the eigenvalues of the Hessenberg matrix

$$\mathbf{H}_n(\lambda) = \begin{bmatrix} \gamma_1 & \gamma_2 & \cdots & \gamma_{n-1} & \beta_n + \lambda \\ \alpha_2 & \gamma_2 & \cdots & \gamma_{n-1} & \beta_n + \lambda \\ 0 & \alpha_3 & \cdots & \gamma_{n-1} & \beta_n + \lambda \\ \vdots & \vdots & & \vdots & \vdots \\ 0 & 0 & \cdots & \gamma_{n-1} & \beta_n + \lambda \\ 0 & 0 & \cdots & \alpha_n & \beta_n + \lambda \end{bmatrix}.$$

The proof of this for $\lambda = 0$ can be found in [7]. Now let

$$A_n(\lambda, z) = \int_a^b \frac{B_n(\lambda, z) - B_n(\lambda, t)}{z - t} d\psi(t), \quad n \geq 1.$$

Then it follows from (2.1) that

$$A_n(\lambda, z) = \int_a^b t^{-p} \frac{t^p B_n(\lambda, z) - z^p B_n(\lambda, t)}{z - t} d\psi(t), \quad n \geq 1, \quad (2.3)$$

where p is any integer such that $0 \leq p \leq n - 1$. Thus, in the relation

$$\frac{A_n(\lambda, z)}{B_n(\lambda, z)} = \sum_{r=1}^n \frac{w_{n,r}^{(\lambda)}}{z - z_{n,r}^{(\lambda)}}, \quad n \geq 1,$$

we have

$$w_{n,r}^{(\lambda)} = \frac{A_n(\lambda, z_{n,r}^{(\lambda)})}{B_n'(\lambda, z_{n,r}^{(\lambda)})} = \frac{\{z_{n,r}^{(\lambda)}\}^p}{B_n'(\lambda, z_{n,r}^{(\lambda)})} \int_a^b \frac{t^{-p} B_n(\lambda, t)}{t - z_{n,r}^{(\lambda)}} d\psi(t), \quad r = 1, 2, \dots, n. \quad (2.4)$$

Now, by considering the interpolation of $t^{n-1} f(t)$ on the zeros of $B_n(\lambda, z)$ we obtain the quadrature rule

$$\int_a^b f(t) d\psi(t) = \sum_{r=1}^n w_{n,r}^{(\lambda)} f(z_{n,r}^{(\lambda)}) + \mathbf{E}_n(\lambda, f), \quad (2.5)$$

where $\mathbf{E}_n(\lambda, f) = 0$ for $t^{n-1} f(t) \in \mathbf{P}_{2n-2}$.

As $\mathbf{E}_n(\lambda, f) = 0$ whenever $f(t) \in \mathbf{P}_{n-1}$, the above is also the usual interpolatory quadrature rule. Taking $f(t) = t^{-n+1} \{B_n(\lambda, t)/(t - z_{n,r}^{(\lambda)})\}^2$, we find that $w_{n,r}^{(\lambda)}$ is positive. Moreover, taking $f(t) = 1$ we get $\sum_{r=1}^n w_{n,r}^{(\lambda)} = \mu_0$.

3. THE $\overline{\text{SCS}}(a, b)$ DISTRIBUTIONS AND $B_n(\lambda, z)$

We assume throughout the rest of this article that $d\psi(t)$ is a $\overline{\text{SCS}}(a, b)$ distribution and that $\lambda \in \Re$ is such that $B_n(\lambda, 0) \neq 0$ (i.e., $\lambda \neq -\beta_n$).

In (2.1), by applying Lemma 1.1 we obtain

$$\int_a^b t^{-n+s} t^n B_n(\lambda, c/t) d\psi(t) = 0, \quad 1 \leq s \leq n-1, n \geq 2.$$

Thus from Theorem 2.1, for each $\lambda \in \Re$ there exists an $\eta \in \Re$ such that

$$\frac{z^n B_n(\lambda, c/z)}{B_n(\lambda, 0)} = B_n(\eta, z), \quad n \geq 1. \quad (3.1)$$

Here we have also included $n = 1$, which is easily seen to hold. Since $B_n(z) = B_n(0, z)$, we have from (1.3)

$$z^n B_n(\lambda, c/z) = z^n B_n(0, c/z) - \lambda z^n B_{n-1}(0, c/z), \quad n \geq 1.$$

Thus from (1.5)

$$\begin{aligned} z^n B_n(\lambda, c/z) &= B_n(0, 0) \{B_n(0, z) - \alpha_{n+1} B_{n-1}(0, z)\} \\ &\quad - \lambda B_{n-1}(0, 0) \{B_n(0, z) + \beta_n B_{n-1}(0, z)\}, \quad n \geq 1. \end{aligned}$$

Regrouping again leads to the relation (3.1), where we now find

$$\eta = \frac{\alpha_{n+1} B_n(0, 0) + \lambda \beta_n B_{n-1}(0, 0)}{B_n(0, 0) - \lambda B_{n-1}(0, 0)}, \quad n \geq 1.$$

Since $B_n(0, 0) = -\beta_n B_{n-1}(0, 0)$, the following result is established.

THEOREM 3.1. *For any $n \geq 1$, if*

$$\eta = \frac{\beta_n(\alpha_{n+1} - \lambda)}{\beta_n + \lambda}, \quad (3.2)$$

then the polynomials $B_n(\lambda, z)$ and $B_n(\eta, z)$ are related to each other by the relation (3.1). Furthermore, when $\lambda, \eta > -\beta_n$,

$$z_{n,r}^{(\lambda)} = c/z_{n,n+1-r}^{(\eta)}, \quad r = 1, 2, \dots, n,$$

and when $\lambda, \eta < -\beta_n$,

$$z_{n,1}^{(\lambda)} = c/z_{n,1}^{(\eta)}, \quad z_{n,r}^{(\lambda)} = c/z_{n,n+2-r}^{(\eta)}, \quad r = 2, 3, \dots, n.$$

The following is a special case of the above theorem.

THEOREM 3.2. For any $n \geq 1$, if λ is equal to

$$\hat{\lambda}_n = \sqrt{\beta_n}(\sqrt{\gamma_n} - \sqrt{\beta_n}) \quad \text{or} \quad \tilde{\lambda}_n = \sqrt{\beta_n}(-\sqrt{\gamma_n} - \sqrt{\beta_n})$$

then

$$\frac{z^n B_n(\lambda, c/z)}{B_n(\lambda, 0)} = B_n(\lambda, z). \quad (3.3)$$

In particular, $z_{n,r}^{(\hat{\lambda}_n)} = c/z_{n,n+1-r}^{(\hat{\lambda}_n)}$, $r = 1, 2, \dots, n$, and $z_{n,1}^{(\tilde{\lambda}_n)} = -\sqrt{c}$, $z_{n,r}^{(\tilde{\lambda}_n)} = c/z_{n,n+2-r}^{(\tilde{\lambda}_n)}$, $r = 2, 3, \dots, n$.

Proof. In Theorem 3.1 letting $\eta = \lambda$ we get $\lambda^2 + 2\beta_n\lambda - \beta_n\alpha_{n+1} = 0$. We take $\hat{\lambda}_n$ as the positive and $\tilde{\lambda}_n$ as the other solution of this quadratic equation. Since $\tilde{\lambda}_n \leq -\beta_n$, the polynomial $B_n(\tilde{\lambda}_n, z)$ has that one negative zero and (3.3) implies that this zero must be equal to $-\sqrt{c}$. This completes the proof.

EXAMPLE. We consider the shifted log-normal distribution

$$d\psi(t) = \frac{1}{2\kappa\sqrt{\pi}} t^{-1} e^{-(\ln(t)/2\kappa)^2} dt, \quad t \in (0, \infty).$$

This distribution, which has also been considered in [3], is a $\text{ScS}(0, \infty)$ distribution with $c = 1$. For this the coefficients of the recurrence relation (1.2) are

$$\beta_n = q^{1/2}, \quad \alpha_{n+1} = q^{1/2}(q^{-n} - 1), \quad n \geq 1,$$

where $q = e^{-2\kappa^2}$. It is interesting to note that this distribution is also a $\text{ScS}(0, \infty)$ distribution with $c = q$.

It is easily verified that these coefficients satisfy (1.6). By following reasonings similar to those of [2, 5], we obtain

$$B_n(z) = \sum_{r=0}^n (-1)^r q^{-r(n-r)} \begin{Bmatrix} n \\ r \end{Bmatrix}_q q^{r/2} z^{n-r}, \quad n \geq 1,$$

where $\left\{ \begin{smallmatrix} n \\ r \end{smallmatrix} \right\}_q$ are the q -binomial coefficients given by

$$\left\{ \begin{smallmatrix} n \\ 0 \end{smallmatrix} \right\}_q = \left\{ \begin{smallmatrix} n \\ n \end{smallmatrix} \right\}_q = 1, \quad \left\{ \begin{smallmatrix} n \\ r \end{smallmatrix} \right\}_q = \frac{\prod_{k=1}^n (1 - q^k)}{\prod_{k=1}^r (1 - q^k) \prod_{k=1}^{n-r} (1 - q^k)}, \quad 1 \leq r \leq n-1,$$

for $n \geq 0$. We can write

$$B_n(\lambda, z) = z^n + \sum_{r=1}^n (-1)^r q^{-r(n-r)} \left\{ \begin{smallmatrix} n \\ r \end{smallmatrix} \right\}_q \left[1 + \lambda q^{n-r} \frac{(1 - q^r)}{(1 - q^n)} \right] q^{r/2} z^{n-r} \quad n \geq 1.$$

Hence, for $\lambda = \hat{\lambda}_n = q^{1/2}(q^{-n/2} - 1)$, we get

$$B_n(\hat{\lambda}_n, z) = \sum_{r=0}^n (-1)^r q^{-r(n-r)} \left\{ \begin{smallmatrix} n \\ r \end{smallmatrix} \right\}_q \left[\frac{q^{r/2} + q^{(n-r)/2}}{1 + q^{n/2}} \right] z^{n-r}, \quad n \geq 1$$

and for $\lambda = \tilde{\lambda}_n = q^{1/2}(-q^{-n/2} - 1)$, we get

$$B_n(\tilde{\lambda}_n, z) = \sum_{r=0}^n (-1)^r q^{-r(n-r)} \left\{ \begin{smallmatrix} n \\ r \end{smallmatrix} \right\}_q \left[\frac{q^{r/2} - q^{(n-r)/2}}{1 - q^{n/2}} \right] z^{n-r}, \quad n \geq 1.$$

From the symmetry of the coefficients of $B_n(\hat{\lambda}_n, z)$ and $B_n(\tilde{\lambda}_n, z)$ we see that these polynomials satisfy the relation (3.3).

4. THE $\text{ScS}(a, b)$ DISTRIBUTIONS AND $w_{n,r}^{(\lambda)}$

From (3.1) we have

$$B'_n(\lambda, c/z) = \frac{nz}{c} B_n(\lambda, c/z) - \frac{z^{-n+2}}{c} B_n(\lambda, 0) B'_n(\eta, z), \quad n \geq 1. \quad (4.1)$$

Now in (2.3) with $p = 0$, substituting c/u for t and then replacing z by c/z , we obtain

$$A_n(\lambda, c/z) = \frac{z}{c} \int_a^b u \frac{B_n(\lambda, c/z) - B_n(\lambda, c/u)}{u - z} d\psi(u).$$

Hence, from (3.1)

$$A_n(\lambda, c/z) = \frac{z^{-n+1}}{c} B_n(\lambda, 0) \int_a^b u^{-n+1} \frac{u^n B_n(\eta, z) - z^n B_n(\eta, u)}{u - z} d\psi(u).$$

Writing $u^n B_n(\eta, z) - z^n B_n(\eta, u)$ as $zu^{n-1} B_n(\eta, z) - z^n B_n(\eta, u) + u^{n-1}(u - z) B_n(\eta, z)$, immediately gives

$$A_n(\lambda, c/z) = \frac{z^{-n+1}}{c} B_n(\lambda, 0) [\mu_0 B_n(\eta, z) - z A_n(\eta, z)], \quad n \geq 1. \quad (4.2)$$

THEOREM 4.1. *For any $\text{SCS}(a, b)$ distribution $d\psi(t)$, if $\lambda, \eta > -\beta_n$ and they satisfy (3.2), then the weights of the quadrature rule (2.5) satisfy*

$$w_{n,r}^{(\lambda)} = w_{n,n+1-r}^{(\eta)}, \quad r = 1, 2, \dots, n, n \geq 1.$$

In particular,

$$w_{n,r}^{(\hat{\lambda}_n)} = w_{n,n+1-r}^{(\hat{\lambda}_n)}, \quad r = 1, 2, \dots, n, n \geq 1.$$

Proof. In (4.1) and (4.2) substituting $z = z_{n,n+1-r}^{(\eta)}$ (i.e., by Theorem 3.1, $c/z = z_{n,r}^{(\lambda)}$), and then using (2.4) immediately gives the required result.

Similarly, we can also obtain the following.

THEOREM 4.2. *For any $\text{SCS}(a, b)$ distribution $d\psi(t)$, if $\lambda, \eta < -\beta_n$ and they satisfy (3.2), then the weights of the quadrature rule (2.5) satisfy*

$$w_{n,1}^{(\lambda)} = w_{n,1}^{(\eta)}, \quad w_{n,r}^{(\lambda)} = w_{n,n+2-r}^{(\eta)}, \quad r = 2, 3, \dots, n, n \geq 2.$$

In particular,

$$w_{n,r}^{(\hat{\lambda}_n)} = w_{n,n+2-r}^{(\hat{\lambda}_n)}, \quad r = 2, 3, \dots, n, n \geq 1.$$

We now define the step functions $\psi_n(t)$, $n \geq 1$, on the interval (a, b) by

$$\psi_n(t) = \begin{cases} 0 & a < t \leq z_{n,1}^{(\hat{\lambda})}, \\ \sum_{s=1}^r w_{n,s}^{(\hat{\lambda})}, & z_{n,r}^{(\hat{\lambda})} < t \leq z_{n,r+1}^{(\hat{\lambda})}, r = 1, 2, \dots, n-1, \\ \mu_0, & z_{n,n}^{(\hat{\lambda})} < t < b, \end{cases}$$

Then Theorems 3.2 and 4.1 tell us that the distributions given by $d\psi_n(t)$ are all $\text{ScS}(a, b)$ distributions.

EXAMPLE. We consider the distribution

$$d\bar{\psi}(t) = \frac{1 + \sqrt{ab}/t}{\sqrt{b-t}\sqrt{t-a}} dt, \quad t \in (a, b),$$

where $0 < a < b < \infty$. This distribution, which was the principle object of study in [8], is a $\text{ScS}(a, b)$ distribution with $c = ab$.

THEOREM 4.3. For the distribution $d\bar{\psi}(t)$ we have for $n \geq 1$

$$z_{n,n+1-r}^{(\lambda_n)} = (\beta + \alpha v_{n,r}) + \sqrt{(\beta + \alpha v_{n,r})^2 - \beta^2}, \quad z_{n,r}^{(\lambda_n)} = c/z_{n,n+1-r}^{(\lambda_n)},$$

for $r = 1, 2, \dots, [(n+1)/2]$ and

$$w_{n,r}^{(\lambda_n)} = \frac{2\pi}{n}, \quad \text{for } r = 1, 2, \dots, n.$$

Here,

$$v_{n,r} = 1 + \cos((2r-1)\pi/n), \beta = \sqrt{ab}, \text{ and } \alpha = (\sqrt{b} - \sqrt{a})^2/4.$$

Proof. Clearly these results satisfy the required conditions given by Theorems 3.2 and 4.1. First we consider the polynomials $\hat{B}_n(z)$, $n \geq 1$, defined by

$$\int_a^b t^{-n+2} \hat{B}_n(t) \frac{1}{\sqrt{b-t}\sqrt{t-a}} dt = 0, \quad 0 \leq s \leq n-1, n \geq 1. \quad (4.3)$$

The distribution in (4.3) is a $\text{ScS}(a, b)$ distribution. The polynomials $\hat{B}_n(z)$ have been studied in some detail in [6], where it was shown that

$$\frac{z^n \hat{B}_n(ab/z)}{\hat{B}_n(0)} = \hat{B}_n(z), \quad n \geq 1. \quad (4.4)$$

Now in (4.3), if we substitute t by ab/t and then use the above relation, we obtain

$$\int_a^b t^{-n+s} \hat{B}_n(t) \frac{\sqrt{ab}/t}{\sqrt{b-t}\sqrt{t-a}} dt = 0, \quad 1 \leq s \leq n, n \geq 1.$$

Adding this with (4.3) gives

$$\int_a^b t^{-n+2} \hat{B}_n(t) d\bar{\psi}(t) = 0, \quad 1 \leq s \leq n-1, n \geq 1,$$

which together with (4.4) yields

$$\hat{B}_n(z) = B_n(\hat{\lambda}_n, z), \quad n \geq 1.$$

Here, $B_n(\hat{\lambda}_n, z)$ are the polynomials given by Theorem 3.2 in relation to the $\text{ScS}(a, b)$ distribution $d\bar{\psi}(t)$.

The proof of the theorem is completed by using a result given in [9].

REFERENCES

1. T. S. CHIHARA, "An Introduction to Orthogonal Polynomials," Mathematics and its Applications Series, Gordon & Breach, New York, 1978.
2. S. C. COOPER, W. B. JONES, AND W. J. THRON, Asymptotics of orthogonal L-polynomials for log-normal distributions, *Construct. Approx.* **8** (1992), 59–67.
3. A. K. COMMON AND J. H. MCCABE, The symmetric strong moment problem, submitted.
4. W. B. JONES, O. NJÅSTAD, AND W. J. THRON, Two point Padé expansions for a family of analytic functions, *J. Comput. Appl. Math* **9** (1983), 105–123.
5. P. I. PASTRO, Orthogonal polynomials and some q-beta integrals of Ramanujan, *J. Math. Anal. Appl.* **112** (1985), 517–540.
6. A. SRI RANGA, Another quadrature rule of highest algebraic degree of precision, *Numer. Math.*, to appear.
7. A. SRI RANGA AND E. X. L. DE ANDRADE, Zeros of polynomials which satisfy a certain three term recurrence relation, *Commun. Anal. Theory Cont. Fractions* **1** (1992), 61–65.
8. A. SRI RANGA AND E. X. L. DE ANDRADE, A weight function that appears in the limit and certain associated polynomials, submitted.
9. A. SRI RANGA AND M. VEIGA, A Tchebyshev type quadrature rule with some interesting properties, *J. Comput Appl. Math.*, to appear.